#### Ocean Uncertainty Quantification Summer School

## **Intro to Data Assimilation**

Sarah Williamson July 16, 2024

#### Outline

- Introduction and notation
- One-dimensional linear problem: a simple example
  - Extension to multi-dimensional problems
- Data assimilation techniques
  - Sequential methods
  - Variational (smoother) methods

#### What is data assimilation?

*Data assimilation* aims to combine available information (data) with a numerical model while taking into account any possible uncertainties



#### Figure 1: Courtesy of Aneesh Subramanian

# What can one do with data assimilation?

### Weather forecasting



#### Improved ocean state reconstruction



Above: NASA ocean visualization, from https:// www.whoi.edu/know-your-ocean/ocean-topics/ how-the-ocean-works/ocean-circulation/currentsgyres-eddies/.

> Right: NASA ECCO heat content reconstruction, from <a href="https://climate.nasa.gov/vital-signs/ocean-warming/?intent=121">https://climate.nasa.gov/vital-signs/ocean-warming/?intent=121</a>

#### **OCEAN HEAT CONTENT CHANGES SINCE 1992 (NASA)**

Data source: Observations from satellites and various ocean measurement devices, including conductivity-temperature-depth instruments (CTDs), Argo profiling floats, eXpendable BathyThermographs (XBTs), instrumented mooring arrays, and icetethered profilers (ITPs). Credit: NASA ECCO



## Sunspot cycle predictions



Three key pieces to data assimilation:

Data,

July 16, 2024 8 / 54

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# Data, numerical model,

Three key pieces to data assimilation:

# Data, numerical model, uncertainties

# Ingredient 1: Data

#### Where do we get data?



From NASA JPL: https://www.jpl.nasa.gov/news/seal-takes-ocean-heattransport-data-to-new-depths

#### Seals are just one example...



# Ingredient 2: Numerical model

#### Numerical models



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- State consists of coupled variables
- For the ocean, these can be,
  - Temperature, salinity, velocities...

# Ingredient 3: Uncertainties

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- Model parameters

Importantly: data assimilation *takes into account* the model and data uncertainties!

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The best model estimate will be one that's computed with statistical info on sources of uncertainties

# Notation and concepts

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Let's define:

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Generally, these variables are multi-dimensional, but we'll also (briefly) look at a one-dimensional problem

#### A couple definitions from statistics

Let p(x) denote a probability distribution, y = x + n an observation, and let E(.) denote the *expectation value* 

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#### Variance

 $E((n - E(n))^2)$  is the variance of *n*; this is the amount *n* will vary around the mean

#### Example: Least-squares linear estimation

#### Goal:

Estimate unknown quantity x given observations

$$y_1 = x + n_1, \quad y_2 = x + n_2$$

with errors  $n_1, n_2$
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Assumptions:

 $E(n_1) = E(n_2) = 0 \implies$  Unbiased observations  $E(n_1^2) = \sigma_1^2, \quad E(n_2^2) = \sigma_2^2$  $E(n_1n_2) = 0 \implies$  Observation errors are uncorrelated

#### Example: Least-squares linear estimation

A linear approximation would be

 $\tilde{x} = a_1 y_1 + a_2 y_2$ 

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To figure out the weights, we'll minimize a model-data misfit:

$$J(x) = \frac{1}{\sigma_1^2}(x - y_1)^2 + \frac{1}{\sigma_2^2}(x - y_2)^2$$

We begin by taking a derivative:

$$\frac{\partial J}{\partial x} = \frac{1}{\sigma_1^2} \frac{\partial}{\partial x} (x - y_1)^2 + \frac{1}{\sigma_2^2} \frac{\partial}{\partial x} (x - y_2)^2$$

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Next, we set it equal to zero and solve for *x*:

$$\implies x = \frac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2} y_1 + \frac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2} y_2$$

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This tells us precisely how to weight the data points:

$$a_1 = rac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2}, \quad a_2 = rac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2}$$

## A sketch of this estimation



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• This means  $\tilde{x}$  is the estimate that minimizes uncertainty

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Then we want the estimate  $\tilde{x}$  to minimize the loss:

$$J(\mathbf{x}) = (\mathbf{E}\mathbf{x} - \mathbf{y})^T \mathbf{R}^{-1} (\mathbf{E}\mathbf{x} - \mathbf{y})$$

Minimizing J yields  $\tilde{x}$ :

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$$\frac{\partial J}{\partial \mathbf{x}} = \frac{\partial}{\partial x} \left[ \mathbf{x}^T \mathbf{E}^T - \mathbf{y}^T \right] \mathbf{R}^{-1} (\mathbf{E}\mathbf{x} - \mathbf{y}) + (\mathbf{E}\mathbf{x} - \mathbf{y})^T \frac{\partial}{\partial x} \left[ \mathbf{R}^{-1} (\mathbf{E}\mathbf{x} - \mathbf{y}) \right]$$

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Setting this equal to zero and solving for *x* gives:

$$\tilde{\boldsymbol{x}} = (\boldsymbol{E}^{T} \boldsymbol{R}^{-1} \boldsymbol{E})^{-1} \boldsymbol{E}^{T} \boldsymbol{R}^{-1} \boldsymbol{y}$$

The estimate

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The corresponding uncertainty is

$$\boldsymbol{P} = E((\tilde{\boldsymbol{x}} - \boldsymbol{x})(\tilde{\boldsymbol{x}} - \boldsymbol{x})^T)$$

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$$\boldsymbol{P} = E((\tilde{\boldsymbol{x}} - \boldsymbol{x})(\tilde{\boldsymbol{x}} - \boldsymbol{x})^T) = (\boldsymbol{E}^T \boldsymbol{R}^{-1} \boldsymbol{E})^{-1}$$

# Sequential methods: Kalman filtering

#### Note: moving to time-dependent notation

#### We'll let $t \in [0, t_f]$ , $\Delta t$ a timestep. We have:

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$$\mathbf{x}(t + \Delta t) = \mathbf{A}\mathbf{x}(t)$$

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- Observations are now time-dependent y(t)
- In general we'll just add a (t)

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Kalman filtering aims to produce the BLUE  $\tilde{x}(t + \Delta t)$  given both  $\tilde{x}(t)$  and  $y(t + \Delta t)$ 

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Which defines for us

$$\boldsymbol{P}(t + \Delta t, -) := \boldsymbol{A} \boldsymbol{P}(t) \boldsymbol{A}^{T}$$
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#### Luckily, we found what ? was for multi-dimensions:

$$? = (\underline{\boldsymbol{E}}^T \underline{\boldsymbol{R}}^{-1} \underline{\boldsymbol{E}})^{-1} \underline{\boldsymbol{E}}^T \underline{\boldsymbol{R}}^{-1}$$

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$$\underline{\boldsymbol{E}} = \begin{bmatrix} \boldsymbol{I} \\ \boldsymbol{E} \end{bmatrix}, \quad \underline{\boldsymbol{R}} = \begin{bmatrix} \boldsymbol{P}(-) & \boldsymbol{0} \\ \boldsymbol{0} & \boldsymbol{R} \end{bmatrix}$$

### Kalman filtering: the Kalman gain matrix

Putting it all together, the BLUE from the Kalman filter is:

$$\tilde{\mathbf{x}} = \left\{ \begin{bmatrix} \mathbf{I} & \mathbf{E}^T \end{bmatrix} \begin{bmatrix} \mathbf{P}^{-1}(-) & 0 \\ 0 & \mathbf{R}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{I} \\ \mathbf{E} \end{bmatrix} \right\}^{-1} \begin{bmatrix} \mathbf{I} \\ \mathbf{E} \end{bmatrix}^T \begin{bmatrix} \mathbf{P}^{-1}(-) & 0 \\ 0 & \mathbf{R}^{-1} \end{bmatrix} \begin{bmatrix} \tilde{\mathbf{x}}(-) \\ \mathbf{y} \end{bmatrix}$$

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$$= \tilde{\mathbf{x}}(-) + \mathbf{P}(-)\mathbf{E}^{T} \left( \mathbf{E}\mathbf{P}(-)\mathbf{E}^{T} + \mathbf{R} \right)^{-1} \left( \mathbf{y} - \mathbf{E}\tilde{\mathbf{x}}(-) \right)$$

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$$P(t + \Delta t) = (\underline{E}^T \underline{R}^{-1} \underline{E})^{-1}$$
  
=  $P(t + \Delta t, -) - K(t + \Delta) EP(t + \Delta t, -)$ 

## Kalman filter equations

$$\begin{split} \tilde{\mathbf{x}}(t + \Delta t, -) &= \mathbf{A} \tilde{\mathbf{x}}(t) \\ \mathbf{P}(t + \Delta t, -) &= \mathbf{A} \mathbf{P}(t) \mathbf{A}^{T} \\ \mathbf{K}(t + \Delta t) &= \mathbf{P}(t + \Delta t, -) \mathbf{E}^{T} \left( \mathbf{E} \mathbf{P}(t + \Delta t, -) \mathbf{E}^{T} + \mathbf{R}(t + \Delta t) \right)^{-1} \\ \tilde{\mathbf{x}}(t + \Delta t) &= \tilde{\mathbf{x}}(t + \Delta t, -) + \mathbf{K}(t + \Delta t) \left( \mathbf{y}(t + \Delta t) - \mathbf{E} \tilde{\mathbf{x}}(t + \Delta t, -) \right) \\ \mathbf{P}(t + \Delta t) &= \mathbf{P}(t, -) - \mathbf{K}(t + \Delta t) \mathbf{E} \mathbf{P}(t, -) \end{split}$$

# Sequential estimation



### Kalman filter discontinuities

From "Potential artifacts in conservation laws and invariants inferred from sequential state estimation" by Wunsch, **Williamson**, Heimbach



# Variational methods: smoothing problem

Consider the model-data misfit

$$J(\boldsymbol{x}) = \sum_{t=t_0}^{t_f} (\boldsymbol{E}\boldsymbol{x}(t) - \boldsymbol{y}(t))^T \boldsymbol{R}^{-1} (\boldsymbol{E}\boldsymbol{x}(t) - \boldsymbol{y}(t))$$

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Variational methods combine available observations with the numerical model by globally adjusting the model to observations

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### Goal

Utilize Lagrange multipliers to compute

 $\frac{\partial J}{\partial \mathsf{x}(t_0)}$ 

# Minimizing J

• Let  $\mu(t)$  denote the Lagrange multipliers

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• Stationary points of *L* correspond to the minima of the constrained optimization problem!

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#### Finding stationary points of $\mathcal{L}$ : the normal equations

We now take derivatives of

$$\mathcal{L}(\boldsymbol{x},\mu) = J(\boldsymbol{x}) + \sum_{t=t_0+\Delta t}^{t_f} \mu(t)^T (\boldsymbol{x}(t) - \boldsymbol{A}\boldsymbol{x}(t-\Delta t))$$

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# Variational framework (for adjoint method)

• Step o: Define the control variables

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- Step 1: Locate Lagrangian stationary point (i.e. solve the normal equations)
- Step 2: Use computed derivative  $\partial \mathcal{L} / \partial \mathbf{x}(t_0)$  in gradient based optimization

One sticky point in adjoint method

Computing model derivative

$$\frac{\partial \boldsymbol{A}(\boldsymbol{x}(t))}{\partial \boldsymbol{x}(t)}^{T}$$

is difficult

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- One can write this out by hand, but it would be tedious
- Making use of automatic differentiation is ideal

# Sketch of the variational method: improving initial condition



# What about other control variables?

### Variational methods are a little flexible

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- We can opt to define (tune) a different control variable
- In these examples we've had a line x(t) = mt + b
- What about tuning other parameters?

# Variational method sketch: tuning *m*



## Variational method sketch: tuning *m* and *b*



# Sequential vs. variational methods



#### Sequential vs. variational methods

From "Potential artifacts in conservation laws and invariants inferred from sequential state estimation" by Wunsch, **Williamson**, Heimbach



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- Depending on application, different methods warranted:
  - forecasting: sequential methods
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- The solution to the DA problem (either filtering or smoothing) depends on the uncertainties that are provided for the background state and the observational error

# Thank you for listening!